The largest eigenvalue of finite rank deformation of Wigner matrices

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1 The model

$$M_N = X_N + A_N := \frac{1}{\sqrt{N}}W_N + A_N$$

where W_N is a $N \times N$ Hermitian matrix such that $(W_N)_{ii}$, $\sqrt{2}Re((W_N)_{ij})_{i < j}$, $\sqrt{2}Im((W_N)_{ij})_{i < j}$ are iid with common distribution μ . μ is assumed to be symmetric with variance σ^2 and it satisfies a Poincaré inequality. A_N is a deterministic, Hermitian matrix.

Example: $\mu = N(0; \sigma^2), X_N \sim GUE(N, \frac{\sigma^2}{N}).$

- 2 Some known result in the non deformed case $(A_N = 0)$
 - Convergence of the spectral measure $\mu_{X_n} := \frac{1}{N} \sum_i \delta_{\lambda_i(X_N)}$ to the Wigner distribution $\mu_{sc} = \frac{1}{2\pi\sigma^2} \sqrt{4\sigma^2 - x^2} \mathbb{1}_{[-2\sigma,2\sigma]}$.

- Convergence a.s. of $\lambda_{max}(X_N)$ to 2σ (the right endpoint of the support of the limiting distribution)
- Fluctuations (Tracy-Widom, Soshnikov)

$$\sigma^{-1} N^{2/3} \left(\lambda_{max}(X_N) - 2\sigma \right) \xrightarrow{\mathcal{L}} \text{T-W distribution } F_2$$

where the distribution F_2 can be expressed with the Fredholm determinant of an operator associated to the Airy kernel.

3 The deformation

 A_N Hermitian of finite rank r (independent of N) with eigenvalues θ_i of multiplicity k_i ; $\theta_1 > \theta_2 > \ldots > \theta_J$. Convergence of the spectral measure to the semicircular distribution μ_{sc} .

What about the extremal eigenvalues?

1) The Gaussian case (Péché)

Ex: θ_1 with multiplicity 1.

1) si $0 \le \theta_1 < \sigma$, $\sigma^{-1} N^{2/3} \left(\lambda_{max}(M_N) - 2\sigma \right) \xrightarrow{\mathcal{L}} F_2$

2) si
$$\theta_1 = \sigma$$
, $\sigma^{-1} N^{2/3} \left(\lambda_{max}(M_N) - 2\sigma \right) \xrightarrow{\mathcal{L}} F_3$

3) si
$$\theta_1 > \sigma$$
, $N^{1/2} \left(\lambda_{max}(M_N) - \rho_{\theta_1} \right) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \sigma_{\theta_1}^2)$
with $\rho_{\theta_1} = \theta_1 + \frac{\sigma^2}{\theta_1} > 2\sigma$.

2) The non Gaussian case for a particular A_N (Féral-Péché)

 A_N is the deformation defined by $(A_N)_{ij} = \frac{\theta}{N}$, so that r = 1 and $\theta_1 = \theta$.

Same TCL as in the Gaussian case, universality of the fluctuations (independent of μ , the distribution of the entries).

3) The non Gaussian case, A_N general

Theorem 1 a.s. behaviour of the spectrum of M_N . Let $J_{+\sigma}$ (resp. $J_{-\sigma}$) be the number of j's such that $\theta_j > \sigma$ (resp. $\theta_j < -\sigma$). (a) $\forall 1 \leq j \leq J_{+\sigma}, \ \forall 1 \leq i \leq k_j,$ $\lambda_{k_1 + \dots + k_{j-1} + i}(M_N) \longrightarrow \rho_{\theta_j}$ a.s., (b) $\lambda_{k_1 + \dots + k_{J+\sigma} + 1}(M_N) \longrightarrow 2\sigma$ a.s., (c) $\lambda_{k_1 + \dots + k_{J-J-\sigma}}(M_N) \longrightarrow -2\sigma$ a.s., (d) $\forall j \geq J - J_{-\sigma} + 1, \ \forall 1 \leq i \leq k_j,$ $\lambda_{k_1 + \dots + k_{j-1} + i}(M_N) \longrightarrow \rho_{\theta_j}$ a.s.

Remark: Same result as in the sample covariance matrices (Bai-Silverstein, Baik-Silverstein)

Theorem 2 Fluctuations

Let $A_N = \operatorname{diag}(\theta, 0, \dots, 0)$ and assume that $\theta > \sigma$. Then

$$\sqrt{N} \Big(\lambda_{max}(M_N) - \rho_\theta \Big) \xrightarrow{\mathcal{L}} (1 - \frac{\sigma^2}{\theta^2}) \Big\{ \mu * \mathcal{N}(0, v_\theta) \Big\}.$$

where $v_{\theta} = v(\theta, \sigma^2, \int x^4 d\mu(x)).$

 \longrightarrow Non universality of the fluctuations.

4 Elements of Proof of Theorem 1

Step 1 Prove that a.s.

$$Spect(M_N) \subset K_{\sigma}(\theta_1, \dots, \theta_J) + [-\epsilon, +\epsilon]$$
 (1)

for N large, where $K_{\sigma}(\theta_1, \cdots, \theta_J) :=$

$$\left\{\rho_{\theta_J};\cdots;\rho_{\theta_{J-J-\sigma+1}}\right\}\cup\left[-2\sigma;2\sigma\right]\cup\left\{\rho_{\theta_{J+\sigma}};\cdots;\rho_{\theta_1}\right\}.$$

Tool: The Stieltjes transform: for $z \in \mathbb{C} \setminus \mathbb{R}$, define $g_N(z) = \operatorname{tr}_N(G_N(z))$ where $G_N(z) = (zI_N - M_N)^{-1}$ is the resolvent of M_N . We set $h_N(z) = \mathbb{E}[g_N(z)]$.

$$g_N(z) = \int \frac{1}{z - x} d\mu_{M_N}(x); \ h_\sigma(z) = \int \frac{1}{z - x} d\mu_{sc}(x).$$

Aim: Obtain a precise estimate

$$h_{\sigma}(z) - h_N(z) + \frac{1}{N}L_{\sigma}(z) = O(\frac{1}{N^2})$$
 (2)

where L_{σ} is the Stieltjes transform of a distribution η with compact support in K_{σ} .

With the help of the inverse Stieltjes transform,

$$\mathbb{E}[\operatorname{tr}_N(\varphi(M_N))] = \int \varphi(x) d\mu_{sc}(x) + \frac{1}{N} \int \varphi(x) d\eta(x) + O(\frac{1}{N^2}),$$

for φ smooth with compact support; and some variance estimates, we deduce from (2)

$$\operatorname{tr}_N \mathbb{1}_{{}^cK^{\varepsilon}_{\sigma}(\theta_1,\cdots,\theta_J)}(M_N) = O(1/N^{\frac{4}{3}}) \ a.s.$$

and therefore the inclusion of the spectrum (1).

Proof of (2):

1) The Gaussian Case:

• The Gaussian integration by parts formula: $\phi : \mathbb{R} \to \mathbb{C}, \ \xi \ standard \ Gaussian$ $\mathbb{E}(\xi \phi(\xi)) = \mathbb{E}(\phi'(\xi)).$ $\Phi : \mathcal{H}_n(\mathbb{C}) \to \mathbb{C}, \ H \in \mathcal{H}_n(\mathbb{C}),$ $\frac{N}{\sigma^2} \mathbb{E}[\operatorname{Tr}(X_N H) \Phi(X_N)] = \mathbb{E}[\Phi'(X_N) \cdot H]$

Apply it for $\Phi(X_N) = [(zI_N - X_N - A_N)^{-1}]_{kl} = G_N(z)_{kl}$ and $H = E_{kl}$; then sum over k and l.

$$\to \sigma^2 \mathbb{E}[g_N^2(z)] - z \mathbb{E}[g_N(z)] + 1 + \frac{1}{N} \mathbb{E}[\operatorname{Tr}(G_N(z)A_N)] = 0$$

$$\rightarrow \sigma^2 h_N^2(z) - zh_N(z) + 1 + \frac{1}{N} \mathbb{E}[\operatorname{Tr}(G_N(z)A_N)] = O(\frac{1}{N^2})$$

Recall that $\sigma^2 h_\sigma^2(z) - zh_\sigma(z) + 1 = 0.$

Estimate for $\mathbb{E}[\operatorname{Tr}(G_N(z)A_N)]$:

 $A_N = U^* \Lambda U$ where Λ is a diagonal matrix with entries $\lambda_i \neq 0$ for $i \leq r, \lambda_i = 0, i > r$. We can show using

- The Gaussian integration by parts formula
- Some variance estimates
- $h_N(z) = h_\sigma(z) + O(\frac{1}{N})$

the estimate

$$\mathbb{E}[\mathrm{Tr}(G_N(z)A_N)] = \sum_{i=1}^r \frac{\lambda_i}{z - \lambda_i - \sigma^2 h_\sigma(z)} + O(\frac{1}{N})$$

Set

$$R_G^{A_N}(z) = \sum_{i=1}^r \frac{\lambda_i}{z - \lambda_i - \sigma^2 h_\sigma(z)} = \sum_{\theta_i \neq 0} k_i \frac{\theta_i}{z - \theta_i - \sigma^2 h_\sigma(z)}.$$

Then,

$$\sigma^2 h_N^2(z) - z h_N(z) + 1 + \frac{1}{N} R_G^{A_N}(z) = O(\frac{1}{N^2})$$

leading to

$$h_N(z) - h_\sigma(z) + \frac{1}{N}L(z) = O(\frac{1}{N^2})$$

where $L(z) = h_{\sigma}^{-1}(z) \mathbb{E}[(z - sc)^{-2}] R_{G}^{A_{N}}(z).$

Question:

- L Stieltjes transform of a distribution ?

- Support of this distribution?

 $\longleftrightarrow \text{Analyticity of } L \text{ (+ conditions); set of singular points.}$ If $|\theta_i| > \sigma, x \in \mathbb{R} \setminus [-2\sigma, 2\sigma],$

$$x - \theta_i - \sigma^2 h_\sigma(x) = 0 \iff x = \theta_i + \frac{\sigma^2}{\theta_i} := \rho_{\theta_i}.$$

2) The non Gaussian case

GIP replaced by: (Khorunzhy, Khoruzhenko, Pastur)

Lemma 1 Let ξ be a real-valued rv such that $\mathbb{E}(|\xi|^{p+2}) < \infty$. Let $\phi : \mathbb{R} \to \mathbb{C}$ such that the first p+1 derivatives are continuous and bounded. Then,

$$\mathbb{E}(\xi\phi(\xi)) = \sum_{a=0}^{p} \frac{\kappa_{a+1}}{a!} \mathbb{E}(\phi^{(a)}(\xi)) + \epsilon$$

where κ_a are the cumulants of ξ , $|\epsilon| \leq C \sup_t |\phi^{(p+1)}(t)| \mathbb{E}(|\xi|^{p+2})$.

Apply to $\xi = Re((X_N)_{ij}), Im((X_N)_{ij}), (X_N)_{ii}$, the odd cumulants vanish (μ symmetric). One must consider the third derivative of $\Phi = (G_N(z))_{kl}$. One obtains:

$$\sigma^2 h_N^2(z) - zh_N(z) + 1 + \frac{1}{N}R(z) = O(\frac{1}{N^2})$$

where $R(z) = R_G^{A_N}(z) + \kappa_4 R_{\Phi'''}^0(z)$.

 $R^0_{\Phi'''}(z)$ is analytic on $\mathbb{C} \setminus [-2\sigma, 2\sigma]$.

Step 2: A.s. convergence of the first extremal eigenvalues of M_N .

Lemma 2 (Weyl) Let B and C be two $N \times N$ Hermitian matrices. For any pair of integers j, k such that $1 \leq j, k \leq N$ and $j + k \leq N + 1$, we have

$$\lambda_{j+k-1}(B+C) \le \lambda_j(B) + \lambda_k(C).$$

For any pair of integers j, k such that $1 \leq j, k \leq N$ and $j + k \geq N + 1$, we have

$$\lambda_j(B) + \lambda_k(C) \le \lambda_{j+k-N}(B+C).$$

5 Sketch of Proof of Theorem 2

$$\lambda_1(M_N) = \theta + \frac{(W_N)_{11}}{\sqrt{N}} + \check{M}_{\cdot 1}^* \widehat{G}(\lambda_1(M_N))\check{M}_{\cdot 1}$$

where \widehat{G} is the resolvent associated to the Wigner matrix of size N - 1 obtained from M_N by removing the first row and column;

$$\check{M}_{.1} = {}^{t} \left((M_N)_{21}, \dots, (M_N)_{N1} \right) = \frac{1}{\sqrt{N}} {}^{t} \left((W_N)_{21}, \dots, (W_N)_{N1} \right).$$

Use the resolvent equation:

$$\widehat{G}(\lambda_1(M_N)) - \widehat{G}(\rho_\theta) = -(\lambda_1(M_N) - \rho_\theta)\widehat{G}(\rho_\theta)\widehat{G}(\lambda_1(M_N))$$

and

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Theorem 3 (Bai-Yao, Baik-Silverstein) Let $B = (b_{ij})$ be a $N \times N$ random Hermitian matrix and $Y_N = (y_1, \ldots, y_N)$ be an independent vector of size N which contains i.i.d standardized entries with bounded fourth moment and such that $\mathbb{E}(y_1^2) = 0$ if y_1 is complex. Assume that

- (i) there exists a constant a > 0 (not depending on N) such that $||B|| \le a$,
- (ii) $\frac{1}{N}$ Tr B^2 converges in probability to a number a_2 ,

(iii) $\frac{1}{N} \sum_{i=1}^{N} b_{ii}^2$ converges in probability to a number a_1^2 .

Then the random variable $\frac{1}{\sqrt{N}}(Y_N^*BY_N - \text{Tr}B)$ converges in distribution to a Gaussian variable with mean zero and variance

$$(\mathbb{E}|y_1|^4 - 1 - t/2)a_1^2 + (t/2)a_2$$

where t = 4 when Y_1 is real and is 2 when y_1 is complex.